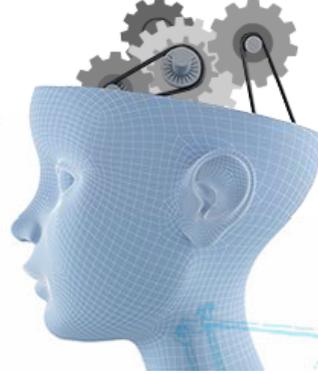


**ENGENHARIA  
MATEMÁTICA**

2º CICLO (MESTRADO)

**FC** FACULDADE DE CIÊNCIAS  
UNIVERSIDADE DO PORTO  
DEPARTAMENTO DE MATEMÁTICA



Apresentação no âmbito da disciplina  
de Seminário de Modelação

# **Regressão Logística na identificação de fatores de risco em recém nascidos prematuros**

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**Título:** Regressão Logística na identificação de fatores de risco em recém nascidos prematuros

- Introdução
- Contextualização & Objetivo
- Artigo
- Método da Regressão Logística
- Trabalho de Investigação a decorrer

**Objetivo:** Determinar fatores de risco em recém nascidos prematuros

- Com o **aumento** da taxa de recém nascidos prematuros, é necessário que haja **intervenção médica**
- Os avanços **médicos** têm salvo algumas crianças prematuras. Contudo, o **risco** de desenvolverem **problemas tardios** é **elevado**

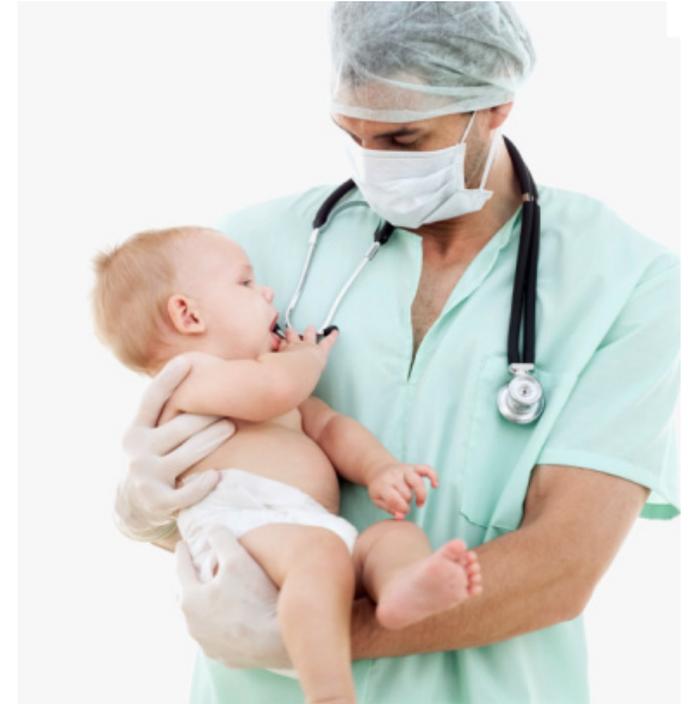
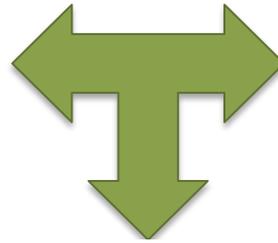


*O que fazer?*

- As estimativas de sobrevivência são difíceis de se obterem, uma vez que, os melhores resultados vêm de **centros de Excelência**

# Contextualização

- A taxa de recém nascidos prematuros tem vindo a **aumentar**
- É necessário **actualizar a informação** destes recém nascidos



- Diminuição da **morbilidade neonatal**

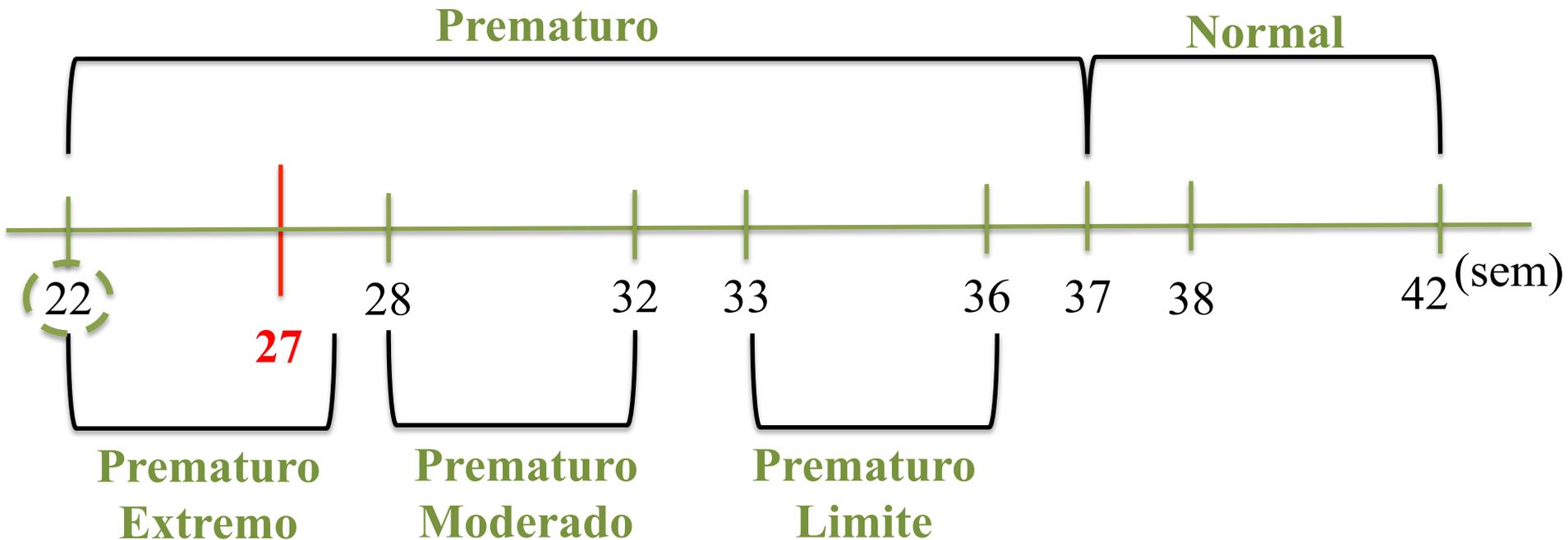


*Diminuição da mortalidade infantil*



# Objetivo

- Fatores **mais precoces** que influenciam a mortalidade a um ano e o **desenvolvimento neurológico** dos recém nascidos prematuros (< 27 semanas de gestação)





## Artigos relacionados com o tema:

- a) I. Seri and J. Evans, “Limits of viability: definition of the gray zone”, *Journal of Perinatology* 2008; 28:S4–S8.
- b) J.R. Kaiser, J.M. Tilford, P.M. Simpson, W.A. Salhab, C.R. Rosenfeld, “Hospital Survival of Very-Low-Birth-Weight Neonates from 1977 to 2000”, *Journal of Perinatology* 2004; 24:343–350.
- c) **The EXPRESS Group Members, “One-Year Survival of Extremely Preterm Infants After Active Perinatal Care in Sweden”, *JAMA* 2009; 301(21):2225–2233.**
- d) J. Peixoto, M. Branco, A. Freitas, C. Dias, “Viabilidade”, *Consensos Nacionais em Neonatologia, Publicação Da secção de Neonatologia Da Sociedade Portuguesa de Pediatria* 2004.

# One-Year Survival of Extremely Preterm Infants After Active Perinatal Care in Sweden

The EXPRESS Group

**T**HE RATE OF PRETERM BIRTHS IS increasing worldwide and the associated neonatal morbidity contributes significantly to infant mortality. Without medical intervention, the number of deaths in those born extremely preterm would equal that of major causes of death in adults.<sup>1</sup> Advances in perinatal medicine have increased survival so that neonatal intensive care can today be life saving even for the most immature infants.<sup>2-19</sup> However, concerns about risks for later disability,<sup>20</sup> adverse developmental programming increasing the risk for hypertension and diabetes in later life, and high costs<sup>21</sup> make intensive care at the limits of viability controversial. Moreover, the evidence for improved outcomes among extremely immature infants, which is fundamental for decision making before, during, and after birth, has been questioned.<sup>22-24</sup> Conversely, misconceptions regarding outcomes may result in suboptimal perinatal care because chances for survival are underestimated.<sup>25</sup>

**Context** Up-to-date information on infant survival after extremely preterm birth is needed for assessing perinatal care services, clinical guidelines, and parental counseling.

**Objective** To determine the 1-year survival in all infants born before 27 gestational weeks in Sweden during 2004-2007.

**Design, Setting, and Patients** Population-based prospective observational study of extremely preterm infants (707 live-born and 304 stillbirths) born to 887 mothers in 904 deliveries (102 multiple births) in all obstetric and neonatal units in Sweden from April 1, 2004, to March 31, 2007.

**Main Outcome Measures** Infant survival to 365 days and survival without major neonatal morbidity (intraventricular hemorrhage grade >2, retinopathy of prematurity stage >2, periventricular leukomalacia, necrotizing enterocolitis, severe bronchopulmonary dysplasia). Associations between perinatal interventions and survival.

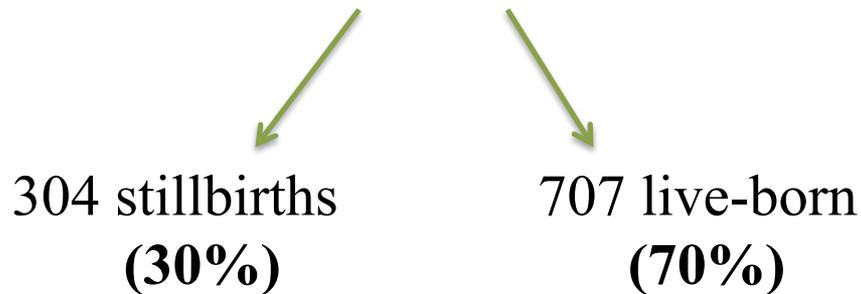
**Results** The incidence of extreme prematurity was 3.3 per 1000 infants. Overall perinatal mortality was 45% (from 93% at 22 weeks to 24% at 26 weeks), with 30% stillbirths, including 6.5% intrapartum deaths. Of live-born infants, 91% were admitted to neonatal intensive care and 70% survived to 1 year of age (95% confidence interval [CI], 67%-73%). The Kaplan-Meier survival estimates for 22, 23, 24, 25, and 26 weeks were 9.8% (95% CI, 4%-23%), 53% (95% CI, 44%-63%), 67% (95% CI, 59%-75%), 82% (95% CI, 76%-87%), and 85% (95% CI, 81%-90%), respectively. Lower risk of infant death was associated with tocolytic treatment (adjusted for gestational age odds ratio [OR], 0.43; 95% CI, 0.36-0.52), antenatal corticosteroids (OR, 0.44; 95% CI, 0.24-0.81), surfactant treatment within 2 hours after birth (OR, 0.47; 95% CI, 0.32-0.71), and birth at a level III hospital (OR, 0.49; 95% CI, 0.32-0.75). Among 1-year survivors, 45% had no major neonatal morbidity.

**Conclusion** During 2004 to 2007, 1-year survival of infants born alive at 22 to 26 weeks of gestation in Sweden was 70% and ranged from 9.8% at 22 weeks to 85% at 26 weeks.

JAMA. 2009;301(21):2225-2233

www.jama.com

- Dados recolhidos na **unidade de obstetrícia e neonatologia**
- Dimensão da amostra : **N=1011** recém nascidos prematuros



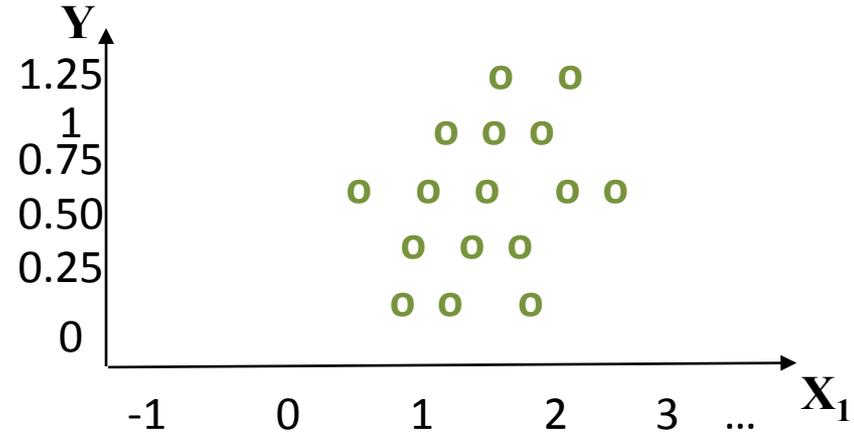
- 877 Mães
- 904 partos dos quais 102 foram **mútiplos**
- **Informação faltante**

- ➔ **Tabela 1:** Características dos Live-born e Intervenções Médicas
- ➔ **Tabela 2:** Mortalidade / Sobrevivência (através de curvas de sobrevivência)
- ➔ **Tabela 3:** Características de Sobrevivência a um ano, com ou sem anomalias
- ➔ **Tabela 4:** Associação entre as intervenções médicas e o risco de mortalidade a um ano para os Live-born [aplicação do método de regressão logística]

- **Modelo de Regressão Linear:**

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_q X_q$$

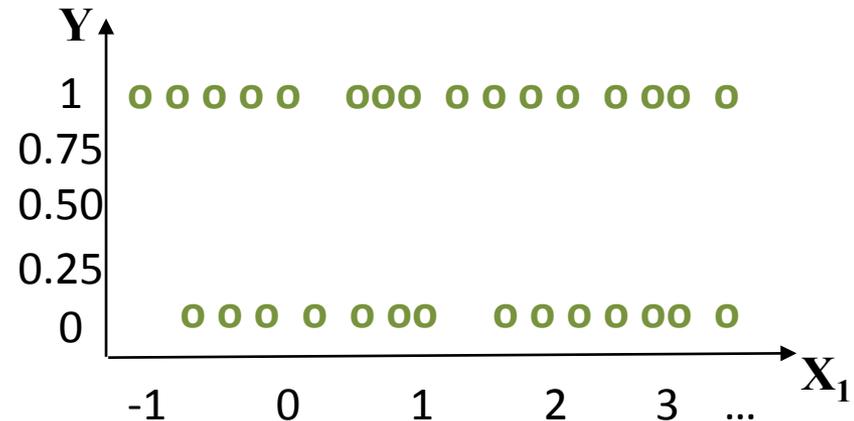
A variável resposta é **contínua**



- **Modelo de Regressão Logística**

$$Y = g(\pi) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_q X_q$$

A variável resposta é **categórica**  
(Neste caso  $Y$  **Binária**)



Para **cada objeto  $i$**  tem-se:

$$Y = g(\pi_i) = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \dots + \beta_q x_{qi}$$

## Modelo Linear Generalizado

**Componente Aleatória**  
⇔ **Y (variável resposta)**

Y tem distribuição  
pertencente à **família  
exponencial**

Amostra dimensão  $n$

**Função de Ligação** ⇔  
função diferenciável e  
monótona  $g$  que associa  
as componentes aleatória  
e sistemática

**Componente Sistemática**  
⇔ combinação linear das  
variáveis explicativas

$$\beta_0 + \beta_1 X_1 + \dots + \beta_q X_q$$

$$Y = g(\pi) = \beta_0 + \beta_1 X_1 + \dots + \beta_q X_q$$

[1] Hosmer, D.W.; Lemeshow, S. (2000). Applied Logistic Regression. 2<sup>nd</sup> ed, John Wiley & Sons, Inc.

**Consideremos [1]:**

- **Variável resposta** :  $Y = \{0,1\}$ , onde 1 é traduzido como o “sucesso”
- $Y_1, \dots, Y_n$  uma **amostra aleatória** de  $Y$
- Vetor de **variáveis explicativas**  $X = (X_1, \dots, X_q)$
- **Observação**  $x_i = (x_1, \dots, x_q)$  de um **objeto i**

$$Y | (X = x_i) \sim \text{Bin}(1, \pi_i)$$

$$\pi_i = \pi(x_i) = P(Y = 1 | X = x_i)$$

$$E(Y | X = x_i) = \pi_i$$

$$\text{Var}(Y | X = x_i) = \pi_i(1 - \pi_i)$$

- Como na regressão linear, assumimos que a **média** pode ser expressa como uma **combinação linear** das variáveis explicativas

$$E(Y | X = x_i) = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \dots + \beta_q x_{qi} \Leftrightarrow$$

$$\Leftrightarrow \pi_i = \beta_0 + \underbrace{\sum_{j=1}^q \beta_j x_{ji}}_{]-\infty, +\infty[}$$

$$0 \leq \pi_i \leq 1$$

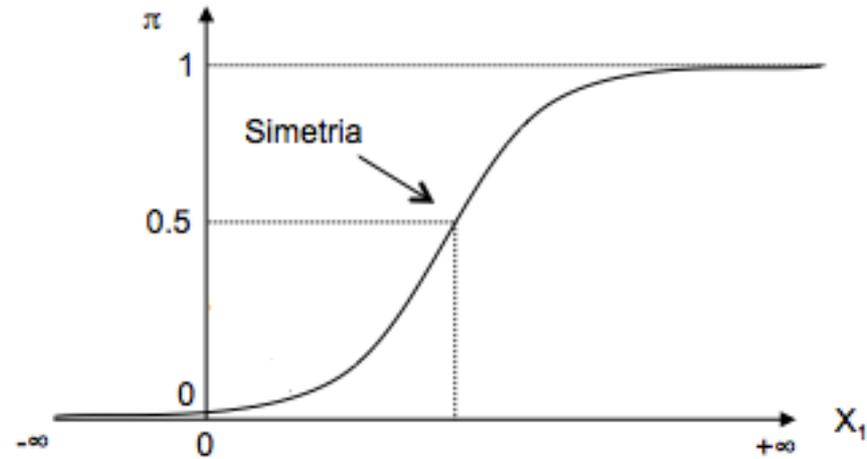
**Coefficientes de regressão**

- **Função de ligação** :  $g(\pi_i) = \text{logit}(\pi_i) = \log\left(\frac{\pi_i}{1 - \pi_i}\right)$

$$\underbrace{\log\left(\frac{\pi_i}{1 - \pi_i}\right)}_{]-\infty, +\infty[} = \beta_0 + \underbrace{\sum_{j=1}^q \beta_j x_{ji}}_{]-\infty, +\infty[}$$

- **Probabilidade de sucesso:**

$$\pi_i = \frac{e^{\beta_0 + \sum_{j=1}^q \beta_j x_{ji}}}{1 + e^{\beta_0 + \sum_{j=1}^q \beta_j x_{ji}}}$$



- **Qual o valor da variável resposta?**

Se  $\pi_i \leq 0.5 \rightarrow Y = 0$

Se  $\pi_i > 0.5 \rightarrow Y = 1$

Nota: A probabilidade de corte **usual** é 0.5. Contudo, existem problemas onde se obtém *melhores resultados* tendo em conta outras probabilidades de corte.

- Os coeficientes da regressão logística ( $\beta$ 's) são estimados através do **método da máxima verosimilhança (MMV)**

- Relembrar que  $Y | (X = x_i) \sim Bin(1, \pi_i)$  e  $f(x_i) = \pi_i^{y_i} (1 - \pi_i)^{1-y_i}$

**Função de Verosimilhança:** 
$$l(\beta; y) = \prod_{i=1}^n f(x_i)$$

- Pelo MMV, as estimativas de  $\beta$ , serão aquelas que maximizem  $l(\beta; y)$

**Função Log – Verosimilhança:** 
$$L(\beta, y) = \ln \left( \prod_{i=1}^n f(x_i) \right) = \sum_{i=1}^n \ln(f(x_i))$$

- Obtém-se as **equações de verosimilhança** para cada  $\beta$
- As estimativas dos coeficientes de regressão obtém-se pelo método iterativo **dos mínimos quadrados pesados**

- **Teste Global de significância das variáveis do modelo**

$\mathbf{H}_0: \beta_1 = \beta_2 = \dots = \beta_q = 0$  ||  $\mathbf{H}_1$ : negar  $\mathbf{H}_0$

$$G = -2 \ln \left[ \frac{\text{verosimilhança do modelo sem as variáveis}}{\text{verosimilhança do modelo com as variáveis}} \right]$$

**Estatística teste:**  $G \sim X^2(q)$ , onde  $q$  é o número de variáveis explicativas

**Rejeitar**  $\mathbf{H}_0$  com nível de significância  $\alpha$  se  $G > X^2_{1-\alpha}(q)$

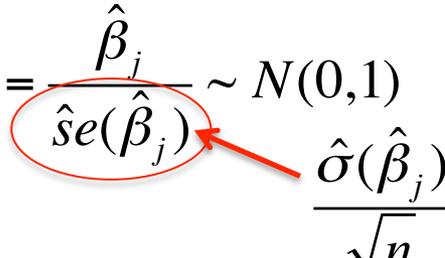
Caso a hipótese nula seja **rejeitada**, avaliar-se-á **individualmente a significância** estatística dos **coeficientes**.



- **Teste de Wald**

$$\mathbf{H}_0: \beta_j = 0, j=1, \dots, q \parallel \mathbf{H}_1: \beta_j \neq 0$$

Estatística teste:  $w_j = \frac{\hat{\beta}_j}{\hat{se}(\hat{\beta}_j)} \sim N(0,1)$



$\hat{\sigma}(\hat{\beta}_j)$   
 $\sqrt{n}$

**Rejeitar  $\mathbf{H}_0$**  com nível de significância  $\alpha$  se  $w_j > N_{1-\alpha}(0,1)$

**Intervalos com  $100(1-\alpha)$  % de confiança para cada coeficiente de regressão,  $\beta_j$  ( $j=1, \dots, q$ )**

$$\hat{\beta}_j \pm z_{\frac{1-\alpha}{2}} \hat{se}(\hat{\beta}_j)$$

**quantil  $(1-\alpha/2)$  da distribuição  $N(0,1)$**

- **Teste da razão de Verosimilhança**

## Comparação de um modelo com o modelo saturado

$H_0$ : o ajustamento do modelo em causa é igual ao ajustamento do modelo saturado ||  $H_1$ : negar  $H_0$

**Desviância** 

$$D = -2 \ln \left[ \frac{\text{verosimilhança do modelo em causa}}{\text{verosimilhança do modelo saturado}} \right] \Leftrightarrow$$
$$\Leftrightarrow D = -2 \sum_{i=1}^n y_i \ln \left( \frac{\hat{\pi}_i}{y_i} \right) + (1 - y_i) \ln \left( \frac{1 - \hat{\pi}_i}{1 - y_i} \right)$$

**Estatística teste:**  $D \sim X^2(p_2 - p_1)$ , onde  $p$  designa o número de parâmetros a estimar em cada modelo

**Rejeitar**  $H_0$  com nível de significância  $\alpha$  se  $D > X^2_{1-\alpha}(p_2 - p_1)$

**H<sub>0</sub>**: os 2 modelos têm a mesma qualidade de ajustamento || **H<sub>1</sub>**: negar H<sub>0</sub>

$$G = -2 \ln \left[ \frac{\text{verossimilhança do modelo sem a variável}}{\text{verossimilhança do modelo com a variável}} \right] \Leftrightarrow$$

$$\Leftrightarrow G = D(\text{verossimilhança do modelo sem a variável}) - D(\text{verossimilhança do modelo com a variável})$$

**Estatística Teste:**  $G \sim X^2(p_2 - p_1)$

**Rejeitar H<sub>0</sub>** com nível de significância  $\alpha$  se  $G > X^2_{1-\alpha}(p_2-p_1)$

- **Teste de Hosmer and Lemeshow**

$H_0$ : O modelo faz um bom ajustamento (não há diferenças entre os valores observados e os valores previstos) ||  $H_1$ : negar  $H_0$

- As n probabilidades de sucesso são ordenadas **crescentemente**
- Tabela de contingência 2xg

Decil	Y=1		Y=0		Total
	Obs	Esp	Obs	Esp	
0-0.1	1				
0.1-0.2	2				
0.2-0.3	3				
•	4				
•	5				
•	6				
•	7				
•	8				
•	9				
0.9-1	10				

**Estatística Teste:**

$$\hat{C} \sim X^2_{(g-2)}$$

**Rejeitar**  $H_0$  com nível de significância  $\alpha$  se  $\hat{C} > X^2_{1-\alpha}(g-2)$

- **Abordagem mais usual** em modelos de regressão logística
- Odds é o quociente entre a probabilidade de **ocorrência de um evento** e a probabilidade de **não ocorrência desse mesmo evento**:

$$odds = \frac{P(sucesso)}{1 - P(sucesso)} = \frac{P(Y = 1)}{1 - P(Y = 1)} = \frac{\pi}{1 - \pi}$$

- Odds ratio traduz-se no **quociente entre dois odds**, permitindo comparar as probabilidades de ocorrência / não ocorrência de um evento para dois objetos/grupos (G) diferentes

$$OR = \frac{odds(G)}{odds(\bar{G})} = \frac{\frac{P(sucesso|G)}{1 - P(sucesso|G)}}{\frac{P(sucesso|\bar{G})}{1 - P(sucesso|\bar{G})}}$$

- Para variáveis **dicotômicas** prova-se que  $OR = e^{\beta_j}, j = 1, \dots, q$
- Análogamente aos **IC** para os parâmetros  $\beta_j$  é possível obter os intervalos com **100(1- $\alpha$ ) %** de confiança para cada **OR**

$$\underbrace{\hat{\beta}_j \pm z_{1-\frac{\alpha}{2}} \hat{se}(\hat{\beta}_j)}_{\text{IC para } \beta_j} \quad \rightarrow \quad \underbrace{e^{\hat{\beta}_j \pm z_{1-\frac{\alpha}{2}} \hat{se}(\hat{\beta}_j)}}_{\text{IC para OR}}$$

**Exemplo:** SGA- Small for gestational age

$$OR[SGA] = \frac{\frac{P(\text{morrer} | SGA)}{P(\text{sobreviver} | SGA)}}{\frac{P(\text{morrer} | \overline{SGA})}{P(\text{sobreviver} | \overline{SGA})}} = 1.47 \quad \text{IC a 95\% [1.22, 1.75], } p\text{-value} < .001$$

**Table 1.** Perinatal Characteristics of Live-Born Infants

Characteristic	No. With Event/No. in Group (%) by Gestational Age in Weeks <sup>a</sup>					Total <27 wk (n = 707)
	≤22 (n = 51) <sup>b</sup>	23 (n = 101)	24 (n = 144)	25 (n = 205)	26 (n = 206) <sup>c</sup>	
Male sex	27/51 (53)	55/101 (55)	78/144 (54)	120/205 (59)	108/206 (52)	388/707 (55)
Twins and triplets	21/51 (41)	16/101 (16)	32/144 (22)	38/205 (18)	51/206 (25)	158/707 (22)
1-min Apgar score ≤3	35/51 (69)	53/101 (53)	63/144 (44)	59/205 (29)	34/206 (17)	244/707 (35)
5-min Apgar score ≤3	39/51 (77)	47/101 (47)	47/144 (33)	42/205 (21)	27/206 (13)	202/707 (29)
Small for gestational age <sup>d</sup>	4/50 (8)	7/100 (7)	16/142 (11)	39/205 (19)	48/206 (23)	114/703 (16)
Birth weight, median (range), g <sup>e</sup>	508 (280 to 730)	590 (320 to 808)	674 (374 to 1070)	784 (266 to 1235)	920 (430 to 1500)	730 (266 to 1500)
Birth weight standard deviation score, median (interquartile range) <sup>f</sup>	-0.07 (-0.88 to -0.51)	-0.44 (-0.60 to 0.19)	-0.59 (-1.36 to 0.02)	-0.76 (-1.50 to -0.02)	-0.87 (-1.87 to -0.16)	-0.67 (-1.47 to 0.06)
Obstetric intervention						
Iatrogenic preterm delivery	4/49 (8)	8/99 (8)	27/137 (20)	59/201 (29)	74/196 (38)	172/682 (25)
Tocolytic treatment <sup>g</sup>	28/45 (62)	71/90 (78)	93/110 (84)	117/142 (82)	104/122 (85)	413/509 (81)
Antenatal steroids	20/49 (40)	85/99 (85)	130/137 (95)	176/198 (89)	180/193 (93)	591/676 (87)
Cesarean delivery	3/51 (6)	17/101 (16)	67/144 (46)	128/205 (62)	141/206 (68)	356/707 (50)
Delivery at level III hospital	24/51 (47)	79/101 (78)	129/144 (90)	166/205 (81)	160/206 (78)	558/707 (79)
Neonatal intervention						
Neonatologist attending at birth	24/50 (48)	83/100 (83)	130/142 (92)	180/205 (88)	170/206 (83)	587/703 (83)
Intubation at birth	13/22 (59)	68/84 (81)	113/142 (80)	126/202 (62)	80/206 (39)	400/656 (61)
Surfactant administration within 2 hours after birth	14/43 (33)	72/95 (76)	120/140 (86)	153/203 (75)	138/205 (67)	497/686 (72)
Admission for neonatal care						
All live-born infants	19/50 (38)	81/100 (81)	132/142 (93)	200/205 (98)	206/206 (100)	638/703 (91)
Outborn infants <sup>h</sup>	6/19 (32)	13/81 (16)	15/132 (11)	37/200 (19)	46/206 (22)	117/638 (18)
Transport to level III care <sup>i</sup>	6/51 (12)	13/101 (13)	15/144 (10)	37/205 (18)	46/206 (22)	117/707 (17)

<sup>a</sup> Number in group denotes the number of infants with available relevant information.

<sup>b</sup> Category denotes gestational age of 22 weeks 0 days to 22 weeks 6 days but also includes 1 infant born at 21 weeks 5 days and 1 infant born at 21 weeks 6 days.

<sup>c</sup> Category denotes gestational age of 26 weeks 0 days to 26 weeks 6 days.

<sup>d</sup> Birth weight smaller than 2 SDs below the mean of the Swedish standard for intrauterine growth.<sup>33</sup>

<sup>e</sup> Reported numbers of infants by gestational age for 22 weeks or younger, 23 weeks, 24 weeks, 25 weeks, and 26 weeks are 50, 100, 142, 205, and 206, respectively.

<sup>f</sup> Standard deviation scores (mean minus actual value/standard deviation) are based on the Swedish standard for intrauterine growth.

<sup>g</sup> Tocolytic treatment in the subgroup with spontaneous preterm labor only.

<sup>h</sup> Percent outborn infants of all live-born infants admitted to level III hospitals.

<sup>i</sup> Infants born at level I or level II hospitals and transported postnatally to level III care as proportion of all admitted live-born infants.

No. With Event/No. in Group (%) by Gestational Age in Weeks<sup>a</sup>

Characteristic	≤22 (n = 51) <sup>b</sup>	23 (n = 101)	24 (n = 144)	25 (n = 205)	26 (n = 206) <sup>c</sup>	Total <27 wk (n = 707)
Obstetric intervention						
iatrogenic preterm delivery	4/49 (8)	8/99 (8)	27/137 (20)	59/201 (29)	74/196 (38)	172/682 (25)
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Transport to level III care <sup>i</sup>	6/51 (12)	13/101 (13)	15/144 (10)	37/205 (18)	46/206 (22)	117/707 (17)

- Existe tendência para o uso de **Tocolytic treatment** , **Antenatal steroids** e **Surfactant** a partir da 23<sup>a</sup> semana gestacional
- Metade dos partos foram **cesariana**
- Dos 707 live-born, 558 nasceram num **hospital de nível III**
- **As intervenções médicas dependem da Idade Gestacional**

**Table 4.** Association Between **Perinatal Interventions** and Risk of Death to Age 1 Year Among Live-Born Infants

		Live-Born Infants		Odds Ratio (95% Confidence Interval) <sup>a</sup>		
		All (n = 707), No. <sup>b</sup>	Dead at 0-364 d (n = 210), No. (%)	Crude	Adjusted <sup>c</sup>	Multivariate Model <sup>d</sup>
Tocolytic treatment						
Yes	$X_1$	413	110 (27)	0.36 (0.23-0.57)	0.43 (0.36-0.52)	0.60 (0.39-0.94)
No		96	48 (50)	1 [Reference]	1 [Reference]	1 [Reference]
Antenatal corticosteroids						
Yes	$X_2$	591	144 (24)	0.31 (0.18-0.54)	0.44 (0.24-0.81)	0.41 (0.20-0.81)
No		85	55 (65)	1 [Reference]	1 [Reference]	1 [Reference]
Cesarean delivery						
Yes	$X_3$	356	75 (21)	0.43 (0.31-0.60)	0.89 (0.60-1.32)	0.98 (0.62-1.52)
No		351	135 (39)	1 [Reference]	1 [Reference]	1 [Reference]
Surfactant administered within 2 hours after birth						
Yes	$X_4$	497	116 (23)	0.47 (0.33-0.68)	0.47 (0.32-0.71)	0.48 (0.31-0.74)
No		189	74 (39)	1 [Reference]	1 [Reference]	1 [Reference]
Born at level III hospital						
Yes	$X_5$	558	145 (26)	0.45 (0.31-0.66)	0.49 (0.32-0.75)	0.78 (0.45-1.35)
No		149	65 (44)	1 [Reference]	1 [Reference]	1 [Reference]

Valores em falta excluídos do estudo



<sup>a</sup>Odds ratios were obtained from logistic regression analysis (simple and adjusted for gestational age) and from a multivariate model including gestational age and all evaluated interventions listed in the table.  
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$$Y = \text{“Mortalidade”} = \begin{cases} 0 = \text{Não} \\ 1 = \text{Sim} \end{cases} \leftarrow \text{Sucesso}$$

$$Y = \text{logit} \left( \frac{\pi}{1 - \pi} \right) = \left( \frac{P(\text{morrer})}{P(\text{sobreviver})} \right)$$

$$Y = \beta_0 + \beta_j X_j, j = 1, 2, 3, 4, 5$$

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- **Todas as variáveis são estatisticamente significativas → IC não contêm o 1**

$$OR = e^{\beta_j} = \frac{P(\text{morrer} | x_j = \text{yes})}{\frac{P(\text{sobreviver} | x_j = \text{yes})}{P(\text{morrer} | x_j = \text{no})}} < 1$$

$$Y = \beta_0 + \beta_j X_j + \beta_{ig} IG, j = 1, 2, 3, 4, 5$$

**Table 4.** Association Between Perinatal Interventions and Risk of Death to Age 1 Year Among Live-Born Infants

		Live-Born Infants		Odds Ratio (95% Confidence Interval) <sup>a</sup>		
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- Parto por **Cesariana** deixou de ser significativo com o ajuste da IG

↓  
Não há diferença entre se realizar cesariana ou não

↓  
Possíveis Medidas Protocolares

# Resultados do artigo

Modelo Simples  
Modelo Ajustado com a IG  
Modelo Multivariado

$$Y = \beta_0 + \sum_{j=1}^5 \beta_j X_j + \beta_{ig} IG$$

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- Parto por **cesariana** continua a não estar associado ao risco de mortalidade, tal como o **nascimento** num **hospital de nível 3**.
- Pelo teste de H&L **não há significância estatística** ( $p\text{-value}=.27$ ) para se rejeitar a hipótese nula de que o modelo faz um bom ajustamento dos dados.

## Título do Projeto:

“ Definição do limite de viabilidade de recém nascidos prematuros usando Análise Discriminante- um estudo na **população portuguesa.**”- Ana Januário

## Parceria entre:



*Serviço  
de Ginecologia  
e Obstetrícia*

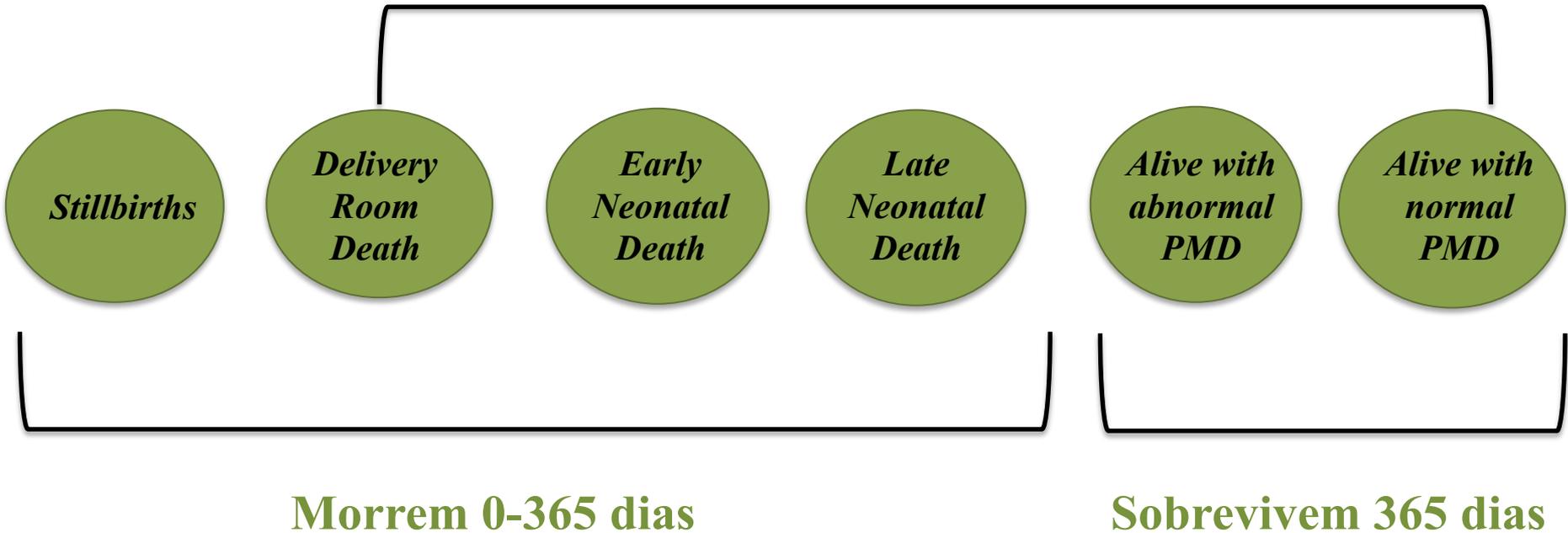
*Serviço  
de  
Neonatologia*

- Dados anónimos de **205** bebés recém nascidos **prematurados** (anteriores às 27 semanas de gestação), seguidos na MJD entre os anos 2000 a 2009
- Dados respeitantes à **mãe**, ao **bebé** e a **intervenções médicas** em diversas fases do estudo

<i>Variável</i>	<i>Legenda</i>	<i>Valor</i>	<i>Valores em falta</i>
<b>IdadeMae</b>	Idade da Mãe(anos)	escala	Tem
<b>GrupoGesta</b>	Primípara	1-sim; 0-não	Tem
<b>GestMult</b>	Gestação Múltipla	1-sim; 0-não	Tem
<b>IG</b>	Idade de gestação (semanas)	escala	Não tem
<b>Peso</b>	Peso(Kg)	escala	Não tem
<b>Sexo</b>	Género	1-masculino;0-feminino	Não tem
<b>DBP</b>	Displasia bronco pulmonar	1-sim; 0-não	Tem
<b>Vigilancia</b>	Gravidez vigiada	1-sim; 0-não	Tem
<b>Corticoides</b>	Corticoides	1-sim; 0-não	Não tem
<b>PartoVaginal</b>	Parto Vaginal	1-sim; 0-não	Tem
.....	.....	.....	.....

- A base de dados possui **valores em falta**

- Duas variáveis relevantes: **Vivo1ano e DPM**

**Live-born**

- **Abordagem** do modelo de **regressão logística** contemplando um maior número de técnicas estatísticas (testes de hipótese, ajustamento, etc)
- Pretende-se **substituir** os **valores em falta** presentes na base de dados

# Risk factors associated with one-year mortality of Extremely premature newborns in Portugal

**A. Januário<sup>1,2</sup>, S. Gouveia<sup>2</sup>, J. Pinto da Costa<sup>1,2</sup>, M. I. Sá<sup>3</sup>, A. Almeida<sup>3</sup>,  
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IJUP'<sup>12</sup>

**5º ENCONTRO  
DE INVESTIGAÇÃO JOVEM  
DA U.PORTO**

# Risk factors associated with one-year mortality of Extremely premature newborns in Portugal

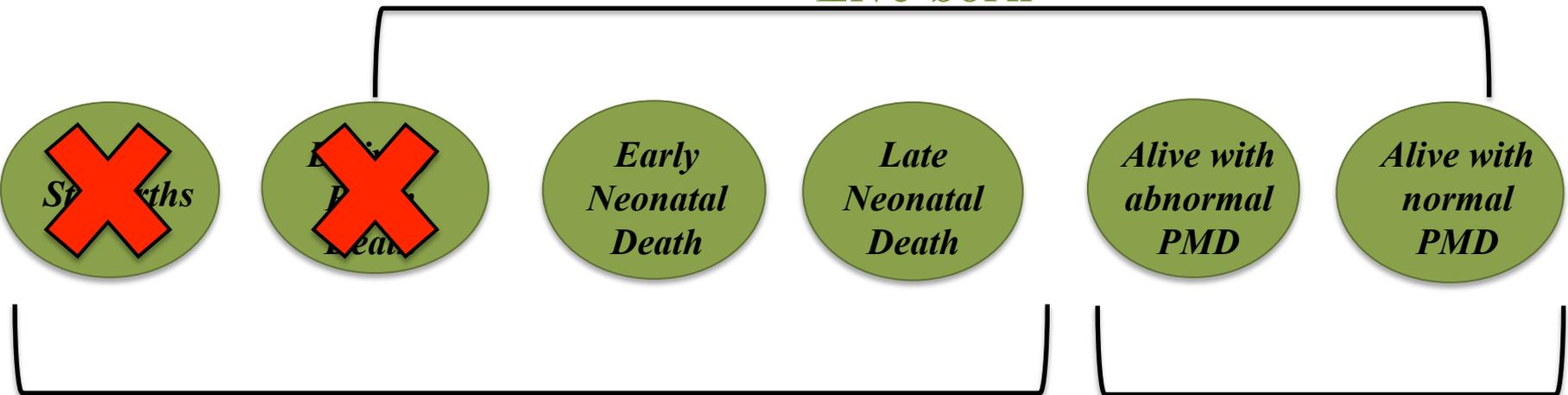
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## Live-born



Morrem 0-365 dias

Sobrevivem 365 dias

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**Table 1:** Mortality odds ratio (OR) and 95% Confidence Intervals for each protocol related factor, where OR<1 indicates decrease risk of mortality.

	All n=167*	Dead n=87	(a) Crude	(b) Adjusted	(c) Multivariate
Pregnancy Surveillance	155	79	0.52(0.15,1.8)	1.12(0.27,4.6)	-----
Antenatal Steroids	142	68	0.29(0.11,0.77)	0.42(0.14,1.24)	-----
Vaginal Delivery	71	45	2.23(1.19,4.17)	1.15(0.55,2.39)	-----
Iatrogenic Delivery	20	10	0.91(0.36,2.31)	1.13(0.4,3.16)	-----
Intubation	119	68	2.04(1.03,4.03)	2.04(0.96,4.33)	2.27(1.04,4.99)
Surfactant	150	81	2.15(0.76,6.12)	2.34(0.75,7.28)	-----
Birth at MJD	146	70	0.22(0.07,0.68)	0.23(0.07,0.77)	0.21(0.06,0.71)
Epoch (2007-2009)	36	21	1.22(0.53,2.82)	1.17(0.45,3)	
(2003-2006)	73	35	0.8(0.4,1.6)	0.85(0.4,1.81)	-----
(2000-2002)	58	31	1[Reference]	1[Reference]	

Binary Logistic Regression: (a) simple; (b) adjusted for GA; (c) multivariate with stepwise procedure including GA; \*=205 cases excluding stillbirths, delivery room death and missing values.

- [1] Hosmer, D.W.; Lemeshow, S. (2000). Applied Logistic Regression. 2<sup>nd</sup> ed, John Wiley & Sons, Inc.
- [2] Fahrmeir, L.; Tutz, G. (2001). Multivariate Statistical Modeling Based on Generalized Linear Models. Second Edition, New York: Springer.
- [3] Agresti, A. (1996). An Introduction to Categorical Data Analysis. New York: John Wiley & Sons, Inc.
- [4] Bastos, J.; Rocha C. (2007). Notas Metodológicas. Análise de Sobrevivência: Métodos Não Paramétricos, pp. 111-114

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